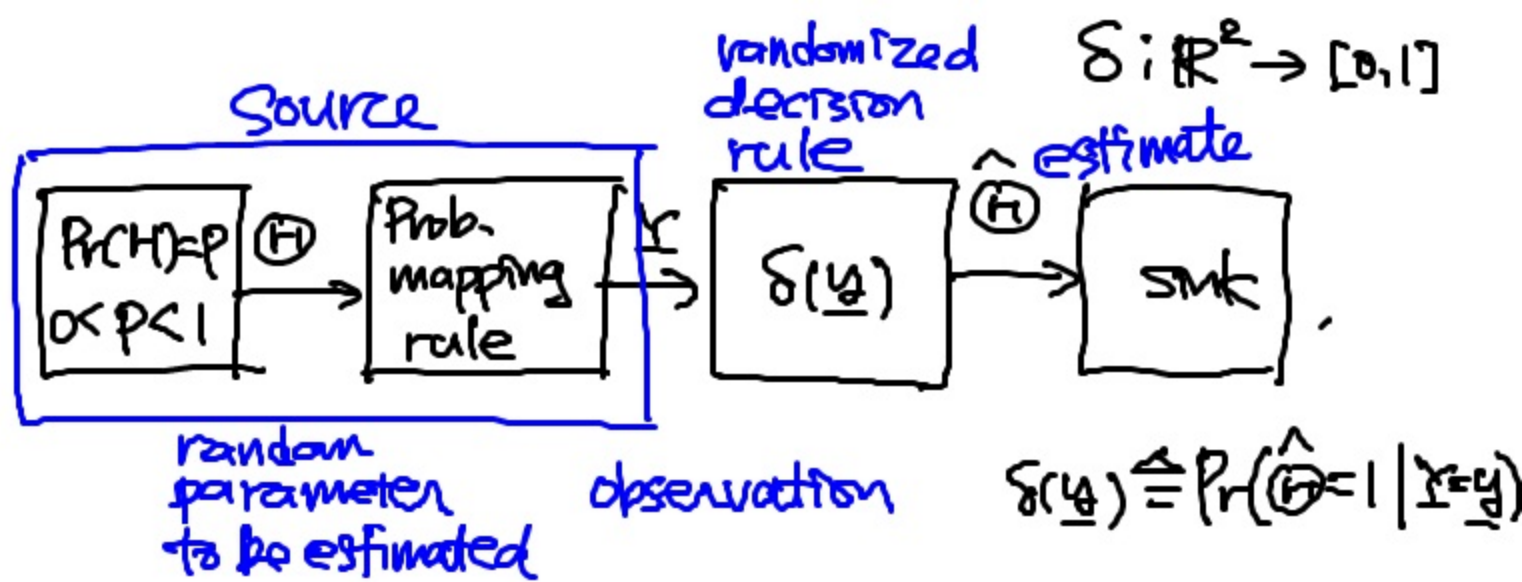


Lec. 03

- Review Given



find

$$\min_{\delta(\underline{y})} \Pr(\theta \neq \hat{\theta})$$

s.t. $0 \leq \delta(\underline{y}) \leq 1, \forall \underline{y} \in \mathbb{R}^2$

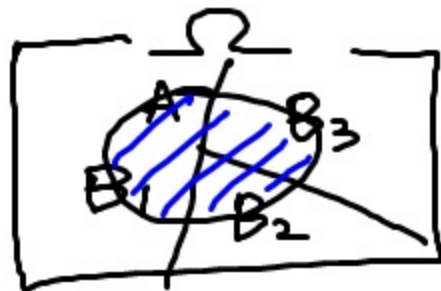
where

$$f_{\underline{Y}}(\underline{y} | \theta = \bar{z}) = \frac{1}{(\sqrt{2\pi}\sigma^2)^2} \exp\left(-\frac{\|\underline{y} - \underline{z}\|^2}{2\sigma^2}\right)$$

$\bar{z} = 1, 2.$

p.2

• Total Probability Theorem



$$P(A) = \sum_i P(A|B_i) P(B_i)$$

$$P(A \cap B_i) = P(A|B_i) P(B_i)$$

where $B_i \cap B_j = \emptyset, \forall i \neq j$ & $\bigcup_i B_i = \Omega$

($\bigcup_i B_i \supset A$ is
also sufficient.)

P.3 $\cdot P(H \neq \hat{H}) = \text{Prob. of error}$
 $x^2 - 2x + 3$
 $= (x-1)^2 + 2$

Representation
(Rewriting)

$$\begin{aligned} D & \left\{ \begin{aligned} & \{H=1\}, \{H=2\} \dots (i) \\ & \{\hat{H}=1\}, \{\hat{H}=2\} \dots (ii) \\ & \{Y=y\} \quad y \in \mathbb{R}^2 \dots (iii) \end{aligned} \right. \end{aligned}$$

(i) $\underbrace{P(H \neq \hat{H} | H=1)}_{\text{conditional prob. of error}} \underbrace{P(H=1)}_{\text{capable of exhaustive search.}} = \text{Intelligence.}$

$+ P(H \neq \hat{H} | H=0) \underbrace{P(H=0)}_{\text{a priori probabilities}}$

$\boxed{5} + \boxed{5} + \boxed{5} + \boxed{5} = 4 + 4 + 4 + 4$

Interpretation.

a priori probabilities
the prior

$= 4 \times 5$
 $= 5 \times 4$

$P^{\dagger} \cdot P(\Theta \neq \hat{\Theta}) = \int_{\underline{y} \in \mathbb{R}^2} P(\Theta \neq \hat{\Theta} | Y = \underline{y}) f_Y(\underline{y}) d\underline{y}$

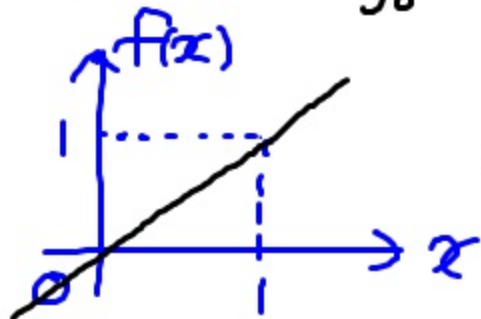
$$f_Y(\underline{y}) = f_{Y|\Theta=1}(\underline{y}) P(\Theta=1) + f_{Y|\Theta=0}(\underline{y}) P(\Theta=0)$$


Q. If $0 \leq f(x) \leq 1, \forall x \in [0, 1]$

then find $f(x)$ s.t.

is maximized.

$$\int_0^1 |f(x) - x| dx$$



$$f(x) = \begin{cases} 1, & 0 \leq x \leq \frac{1}{2} \\ 0, & \frac{1}{2} \leq x \leq 1 \end{cases}$$

p5

$$\min_{\delta(\underline{y})} \Pr(\Theta \neq \hat{\Theta})$$

s.t. $0 \leq \delta(\underline{y}) \leq 1, \forall \underline{y} \in \mathbb{R}^2$

$$\equiv \left(\begin{array}{l} \min \Pr(\Theta \neq \hat{\Theta} | \underline{Y} = \underline{y}) \\ \delta(\underline{y}) \\ \text{s.t. } 0 \leq \delta(\underline{y}) \leq 1 \end{array} \right)_{\underline{y} \in \mathbb{R}^2}$$

equiv

$$\{\hat{\Theta} = 0\} \cap \{Y = \underline{y}\} \quad 1 - \delta(\underline{y})$$

$$\Pr(\Theta \neq \hat{\Theta} | \underline{Y} = \underline{y}) = \Pr(\Theta \neq \hat{\Theta} | \hat{\Theta} = 0, \underline{Y} = \underline{y}) \Pr(\hat{\Theta} = 0 | \underline{Y} = \underline{y})$$

$$+ \Pr(\Theta \neq \hat{\Theta} | \hat{\Theta} = 1, \underline{Y} = \underline{y}) \Pr(\hat{\Theta} = 1 | \underline{Y} = \underline{y})$$

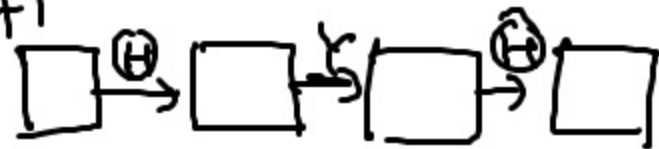
the posterior prob.

$$= \Pr(\Theta = 1 | \underline{Y} = \underline{y}) (1 - \delta(\underline{y})) + \Pr(\Theta = 0 | \underline{Y} = \underline{y}) \delta(\underline{y})$$

a posterior prob.

$$P(A) = \sum_i P(A \cap B_i) = \sum_i P(A | B_i) P(B_i)$$

q1



$\Theta \rightarrow Y \rightarrow \hat{\Theta}$ forms a Markov chain.

p.6

A posteriori prob.

$$\Pr(\Theta=i | Y=y) = \frac{f_Y|\Theta=i(y) \Pr(\Theta=i)}{\sum_{j=0}^1 f_Y|\Theta=j(y) \Pr(\Theta=j)}$$

The Bayes theorem (events version)

$$\Pr(B_i|A) = \frac{\Pr(B_i \cap A)}{\Pr(A)} = \frac{\Pr(A|B_i) \Pr(B_i)}{\sum_j \Pr(A|B_j) \Pr(B_j)}$$

where $(B_i)_i$ forms a partition of Ω .

$$\begin{aligned} - \min_{\delta(y)} \Pr(\Theta \neq \hat{\Theta}) &\equiv \min_{\delta(y)} \Pr(\Theta \neq \hat{\Theta} | Y=y) \\ &\equiv \min_{\delta(y)} \Pr(\Theta=1 | Y=y) (1-\delta(y)) + \Pr(\Theta=0 | Y=y) \delta(y) \end{aligned}$$

$$P.7 \min 5(1-x) + 4x$$

$$0 < x < 1$$

$$\min A(1-\delta) + B\delta$$

$$0 < \delta < 1$$

If $B > A$, then $\delta_{opt} = 0$

If $B = A$, then
 $0 \leq \delta_{opt} \leq 1$

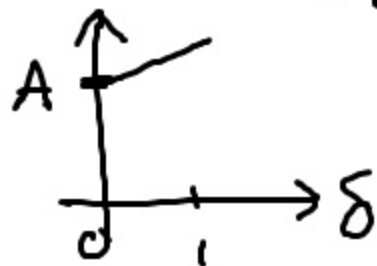
If $B < A$, then
 $\delta_{opt} = 1$

$$5 - 5x + 4x = 5 - x$$



$$A - A\delta + B\delta = A - (A-B)\delta$$

$$= A + (B-A)\delta$$



p.8 ~~min~~ $\Pr(\Theta=0|Y=y) \delta(y) + \Pr(\Theta=1|Y=y) (1-\delta(y))$

$$= \frac{f_{Y|\Theta=0}(y) \Pr(\Theta=0)}{\cancel{f_Y(y)}} \delta(y) + \frac{f_{Y|\Theta=1}(y) \Pr(\Theta=1)}{\cancel{f_Y(y)}} (1-\delta(y))$$

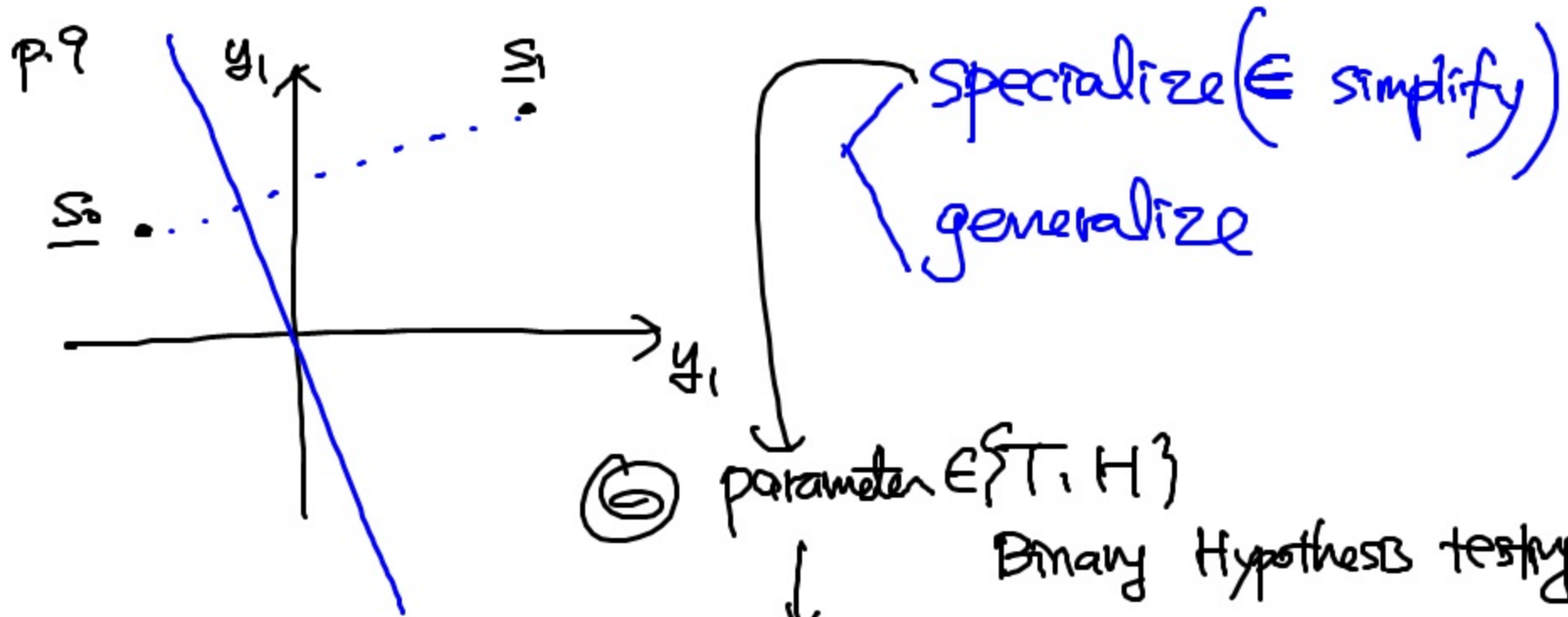
.....

$$\equiv \delta(y) = \begin{cases} 0, & \text{for } \frac{f_{Y|\Theta=1}(y)}{f_{Y|\Theta=0}(y)} \leq \frac{\Pr(\Theta=0)}{\Pr(\Theta=1)} \\ 1, & \text{for } \frac{f_{Y|\Theta=1}(y)}{f_{Y|\Theta=0}(y)} > \frac{\Pr(\Theta=0)}{\Pr(\Theta=1)} \end{cases}$$

LRT

test.

likelihood ratio



⑥ M-ary Hypothesis Testing.