

Lec. 30

Announcement : Quiz #5 next Friday

Corrections : Last lectures results holds only when C does not change as a function of the time index n .

Review

i.i.d. sequence observation : $x_1, x_2, x_3, \dots$

binary hypothesis testing : $\{H=0, H=1\}$

random
cost

$P(H=1) \leq p$
 $C(0,0), C(0,1), C(1,0), C(1,1)$

p.2

objective: To find a sequential detector

$(\delta_n)_{n=0}^{\infty}$ such that

$$\begin{aligned}\delta_n &= \delta_n(\underline{y}_1, \underline{y}_2, \dots, \underline{y}_n) \\ &= (\underbrace{\phi_n(\underline{y}_1, \underline{y}_2, \dots, \underline{y}_n)}_{\text{stopping rule}}, \underbrace{\hat{\theta}_n(\underline{y}_1, \underline{y}_2, \dots, \underline{y}_n)}_{\text{decision rule}})\end{aligned}$$

where $\phi_n = 1 \iff$ Stop at time n

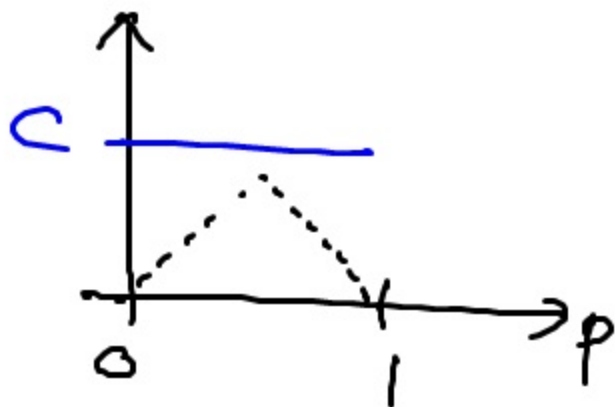
$\hat{\theta}_n = \begin{cases} 0 \\ 1 \end{cases} \iff \begin{array}{l} \text{say } H_0 \\ \text{say } H_1 \end{array}$

p.3

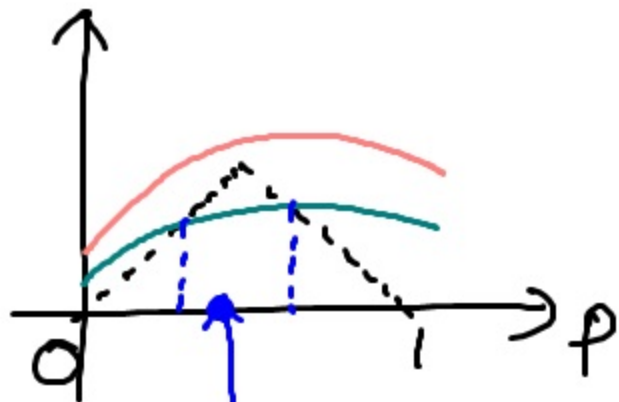
ex/ $C_{\theta, \theta} = 1 - \delta_{\theta, \theta}$ uniform

$$C > \frac{1}{2}$$

$$\delta_0 = (\phi_0, \theta_0) = \begin{cases} (1, 0) \\ (1, 1) \end{cases}$$

ex/ $C_{\theta, \theta} = 1 - \delta_{\theta, \theta}$

$$C > 0$$



$$Pr(\theta=1), \quad Pr(\theta=1 | Y_1=y_1)$$

p4

SPRT(A,B) is optimal for some A, B.

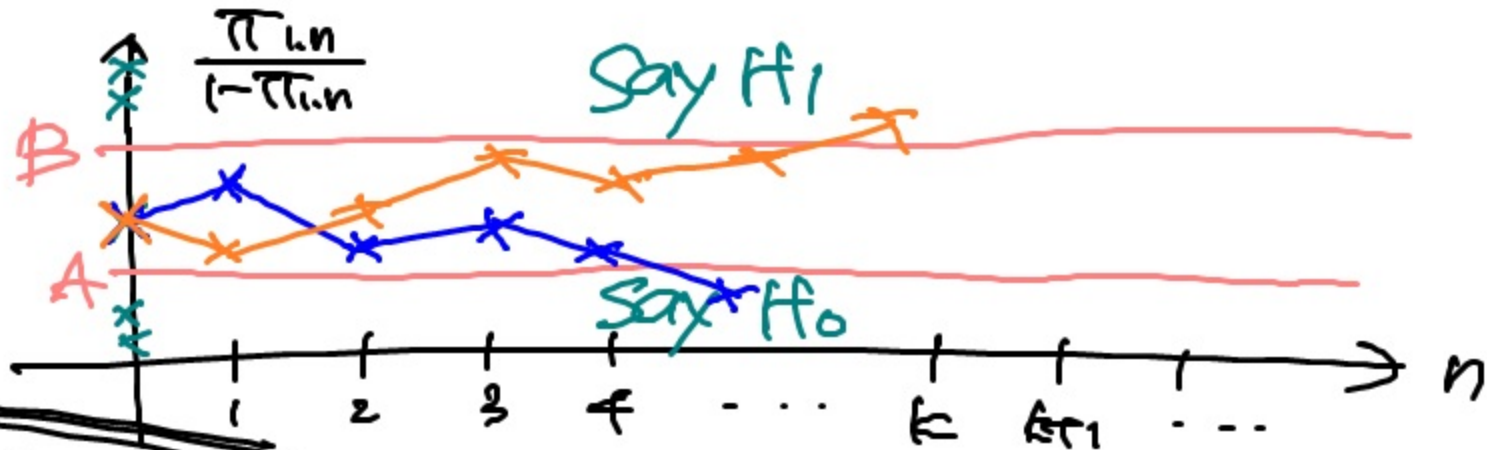
↑↑
fixed!

$$\frac{\Pr(\Theta=1 \mid \underline{y}_1=y_1, \underline{y}_2=y_2, \dots, \underline{y}_n=y_n)}{\Pr(\Theta=0 \mid \underline{y}_1=y_1, \underline{y}_2=y_2, \dots, \underline{y}_n=y_n)}$$

$$\triangleq \frac{\pi_{1,n}}{\pi_{0,n}} = \frac{\pi_{1,n}}{1-\pi_{1,n}} = \prod_{k=1}^n \frac{f(\underline{y}_k \mid \Theta=1)}{f(\underline{y}_k \mid \Theta=0)}$$

$$n=0 \quad \frac{\Pr(\Theta=1)}{\Pr(\Theta=0)}$$

p5



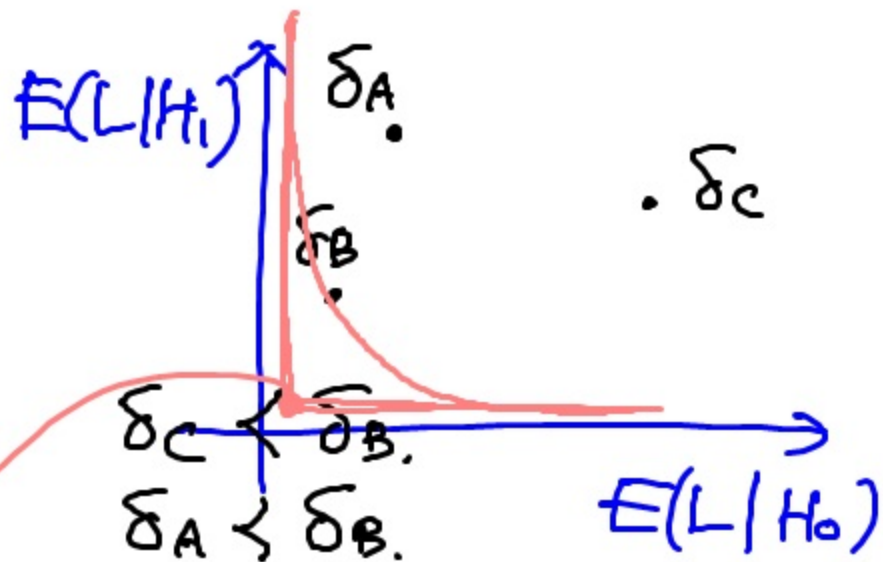
~~Today~~

θ is not random

No δ_0 .

Starts w/ δ_1

SPRT(A, B)



$\phi 6$ $L \triangleq$ Stopping time. Λ_n : prob.-ratio @ time n .

$$P_F = \Pr(\hat{\theta}_L(\underline{Y}_1, \dots, \underline{Y}_L) = 1; \theta = 0)$$

$$= \Pr(\Lambda_L(\underline{Y}_1, \underline{Y}_2, \dots, \underline{Y}_L) \geq B; \theta = 0)$$

$$= \Pr\left(\frac{f_{\underline{Y}_1; \theta=1}(\underline{Y}_1)}{f_{\underline{Y}_1; \theta=0}(\underline{Y}_1)} \times \dots \times \frac{f_{\underline{Y}_L; \theta=1}(\underline{Y}_L)}{f_{\underline{Y}_L; \theta=0}(\underline{Y}_L)} \geq B; \theta = 0\right)$$

$$= \sum_{l=1}^{\infty} \Pr(\Lambda_l(\underline{Y}_1, \dots, \underline{Y}_l) \geq B; \theta = 0)$$

$$= \sum_{l=1}^{\infty} \Pr(\hat{\theta}_l(\underline{Y}_1, \underline{Y}_2, \dots, \underline{Y}_l) = 1; \theta = 0)$$

$$\begin{aligned}
 P^{\Pi} &= P_F \\
 &= P(\hat{\theta}_2(\underline{y}_1, \dots, \underline{y}_L) = 1; \theta = 0) \leq \frac{1}{B} \prod_{n=1}^L f_{\underline{y}_n; \theta=1}(\underline{y}_n) \\
 &= \int_{\underline{y}_1, \underline{y}_2, \dots, \underline{y}_L} \left(\prod_{n=1}^L f_{\underline{y}_n; \theta=0}(\underline{y}_n) \right) \mathbb{1}_{\Lambda_L \geq B}(\underline{y}_1, \dots, \underline{y}_L) d\underline{y}_1 \dots d\underline{y}_L
 \end{aligned}$$

$$\Lambda_L(\underline{y}_1, \dots, \underline{y}_L) = \frac{\prod_{n=1}^L f_{\underline{y}_n; \theta=1}(\underline{y}_n)}{\prod_{n=1}^L f_{\underline{y}_n; \theta=0}(\underline{y}_n)} \geq B$$

$$\leq \frac{1}{B} \int_{\underline{y}_1, \dots, \underline{y}_L} \left(\prod_{n=1}^L f_{\underline{y}_n; \theta=1}(\underline{y}_n) \right) \mathbb{1}_{\Lambda_L \geq B}(\underline{y}_1, \dots, \underline{y}_L) d\underline{y}_1 \dots d\underline{y}_L$$

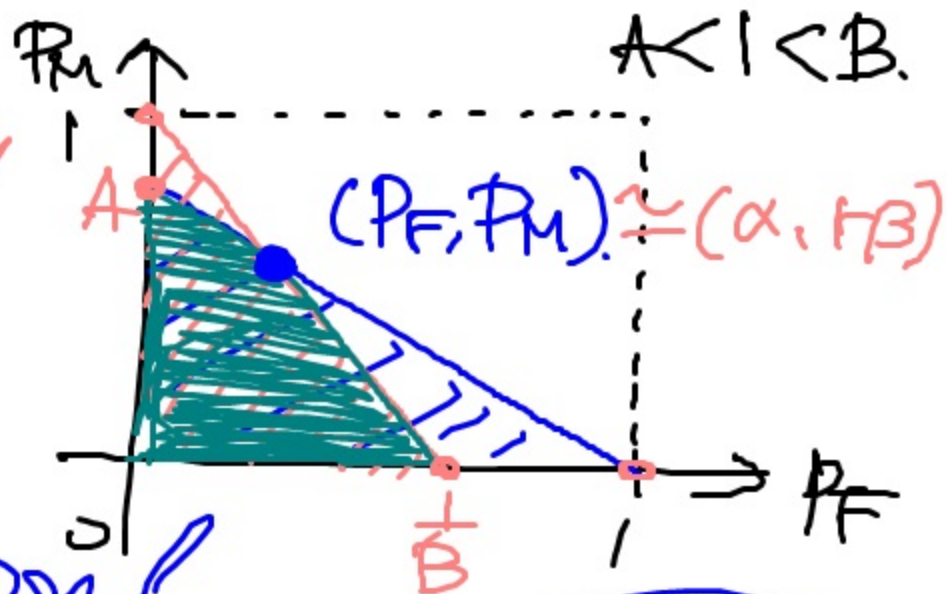
$$= \frac{1}{B} P_0 = \frac{1}{B} (1 - P_M)$$

$$\begin{aligned}
 P_F &\leq \frac{1}{B} (1 - P_M) \\
 P_M &\leq A (1 - P_F)
 \end{aligned}$$

P8

$$P_F \leq \frac{1}{B} (1 - P_M) \quad \text{////}$$

$$P_M \leq A (1 - P_F) \quad \text{////}$$



Wald's approximation!

$$P_F \approx \frac{1}{B} (1 - P_M)$$

$$P_M \approx A (1 - P_F)$$

$$P_F \leq \alpha$$

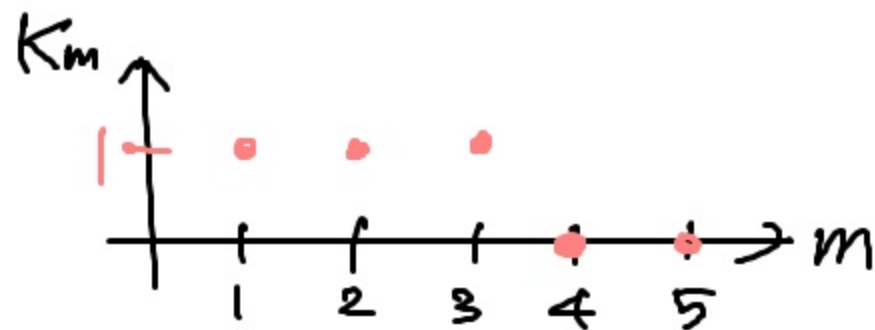
$$P_M \leq 1/B$$

$$A \approx \frac{1 - B}{1 - \alpha}$$

$$B \approx \frac{B}{\alpha}$$

p9 $E(L|H_1), E(L|H_0)$ (L : stopping time.

for $L=3$



K_m : "SPRT(A,B)
observes Y_m !"

$\begin{cases} 1. \\ 0. \end{cases}$

$$\sum_{m=1}^{\infty} K_m = L.$$

" K_1 is not a ft of observation.

$K=1$."

" K_2 is a ft only of Y_1 ."

$K_m(Y_1, \dots, Y_{m-1})$.

$$P^{10} \quad E(L; \theta = \hat{\tau})$$

$$E(\ln L(\underline{Y}_1, \dots, \underline{Y}_L); \theta = \hat{\tau})$$

$$= E\left(\ln \frac{\prod_{l=1}^L f_{Y_l; \theta=1}(\underline{Y}_l)}{\prod_{l=1}^L f_{Y_l; \theta=0}(\underline{Y}_l)}; \theta = \hat{\tau}\right)$$

$$= E\left(\sum_{l=1}^L \ln \frac{f_{Y_l; \theta=1}(\underline{Y}_l)}{f_{Y_l; \theta=0}(\underline{Y}_l)}; \theta = \hat{\tau}\right)$$

$$= E\left(\sum_{m=1}^{\infty} K_m(\underline{Y}_1, \underline{Y}_2, \dots, \underline{Y}_{m-1}) \ln \frac{f_{Y_m; \theta=1}(\underline{Y}_m)}{f_{Y_m; \theta=0}(\underline{Y}_m)}; \theta = \hat{\tau}\right)$$

$p \parallel$

$$= \sum_{m=1}^8 E \left[K_m(\underline{X}_1, \dots, \underline{X}_{m-1}) \ln \frac{f_{\underline{X}_m; \theta=1}(\underline{X}_m)}{f_{\underline{X}_m; \theta=0}(\underline{X}_m)} ; \theta = \hat{\tau} \right]$$

statistically independent

$$= \left(\sum_{m=1}^8 E \left[K_m(\underline{X}_1, \dots, \underline{X}_{m-1}) ; \theta = \hat{\tau} \right] \right) \times E \left[\ln \frac{f_{\underline{X}_m; \theta=1}(\underline{X}_m)}{f_{\underline{X}_m; \theta=0}(\underline{X}_m)} ; \theta = \hat{\tau} \right]$$

$$= E \left[L ; \theta = \hat{\tau} \right] \times E \left[\ln \frac{f_{\underline{X}_m; \theta=1}(\underline{X}_m)}{f_{\underline{X}_m; \theta=0}(\underline{X}_m)} ; \theta = \hat{\tau} \right]$$

$$= E \left[\ln \Lambda_L(\underline{X}_1, \dots, \underline{X}_L) ; \theta = \hat{\tau} \right] \stackrel{\text{Wald's approximation}}{=}$$

P.12

$$E [\ln \Lambda(\mathcal{Y}, \cdot, \mathcal{Z}) ; \theta=0]$$

$$\cong \ln A (1 - P_F) + \ln B \cdot P_F \cong \ln A (1 - \alpha) + \ln B \cdot \alpha$$

$$E [\ln \Lambda(\mathcal{Y}, \cdot, \mathcal{Z}) ; \theta=1]$$

$$\cong \ln A \cdot P_M + \ln B (1 - P_M)$$

$$\cong \ln A (1 - \beta) + \ln B \cdot \beta$$

* SPRT(A,B) stops almost surely.

ΦB

We take another sample if

$$A < \Lambda_e(\underline{y}_1, \dots, \underline{y}_e) < B$$

$$|\Lambda_e(\underline{y}_1, \dots, \underline{y}_e)|$$

