

# □ MS Periodicity

Motivated by  $x(t) = x(t+T), \forall t$

$$\Leftrightarrow x(t) - x(t+T) = 0, \forall t$$

$$\Leftrightarrow |x(t) - x(t+T)|^2 = 0, \forall t$$

Def.  $E\{|x(t) - x(t+T)|^2\} = 0, \forall t$

Lemma.  $x(t) = x(t+T), \forall t$  in the MS sense

$$\Rightarrow \begin{cases} \text{(i)} & \mu_x(t) = \mu_x(t+T), \\ \text{(ii)} & r_{xx}(t, s) = r_{xx}(t+T, s), \forall(t, s) \\ & = r_{xx}(t, s+T), \forall(t, s) \\ & = r_{xx}(t+mT, s+mT), \forall(t, s), \forall(m, n) \end{cases}$$

proof (i) By  $E\{|x|^2\} \geq |E\{x\}|^2$ .

$$\int_{-\infty}^{\infty} x^2 f_x(x) dx \geq \left( \int_{-\infty}^{\infty} x f_x(x) dx \right)^2$$

We have

$$0 \leq |E\{x(t) - x(t+T)\}|^2 \leq E\{|x(t) - x(t+T)|^2\} = 0, \forall t$$

$$\Rightarrow E\{x(t) - x(t+T)\} = 0, \forall t \Leftrightarrow \mu_x(t) = \mu_x(t+T), \forall t$$

$$\text{(ii) By } |E\{x^*y\}|^2 \leq E\{|x|^2\} E\{|y|^2\}.$$

$$\langle \underline{a}, \underline{b} \rangle \leq \sqrt{\langle \underline{a}, \underline{a} \rangle \langle \underline{b}, \underline{b} \rangle}$$

(ii) By  $E\{XY^* \}^2 \leq E\{X\}^2 E\{Y\}^2$ .

we have

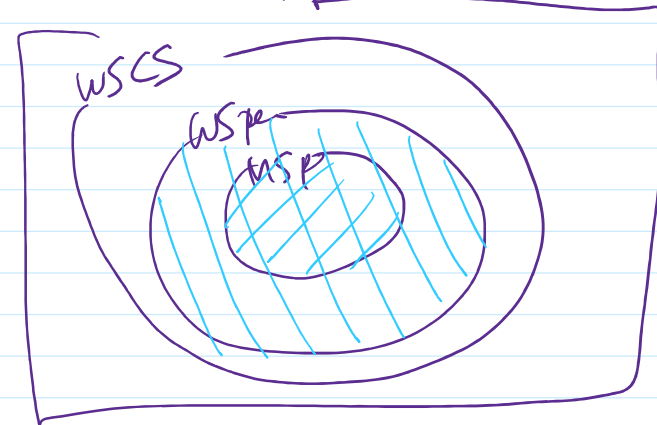
$$0 \leq \underbrace{E\left[X(t) \{X(s) - X(s+T)\}^*\right]^2}_{=0, \forall (t,s)} \leq E\{X(t)\}^2 \underbrace{\left\{E\{X(s) - X(s+T)\}^2\right\}}_{=0, \forall s}, \forall (t,s)$$

$$\Rightarrow E\{X(t)(X(s) - X(s+T))^*\} = 0$$

$$\Leftrightarrow \left. \begin{aligned} r_{xx}(t,s) &= r_{xx}(t, s+T), \forall (t,s) \\ &= r_{xx}(t+T, s), \forall (t,s) \end{aligned} \right\} \Rightarrow$$

$$r_{xx}(t,s) = r_{xx}(t+mT, s+nT), \forall (t,s) \forall (m,n)$$

2nd order RPs



□ Recall Def.  $X(t)$  is WSCS w/ period  $T$  if

(i)  $\mu_X(t) = \mu_X(t+T), \forall t$

(ii)  $r_{xx}(t,s) = r_{xx}(t+T, s+T), \forall (t,s)$

MS periodicity  $\Rightarrow$  WS periodicity  $\Rightarrow$  WSCS

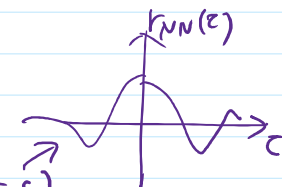
□ cf)



...  $\left\{ \begin{array}{l} X(t) \text{ WSS Gaussian} \\ \leadsto S_{XX}(f) \end{array} \right.$

$$\mu_X(t) = a, \forall t$$

$$r_{XX}(t,s) = r_{XX}(t-s)$$



MS Stationarity

$$E\{|X(t) - X(t-\tau)|^2\} = 0, \forall t, \tau$$

MS periodicity

$$E\{|X(t) - X(t+\tau)|^2\} = 0$$

$$E\{|X(t) - X(t-\tau)|^2\} = 0$$

$$x(t) = x(t-\tau), \forall t, \tau$$

$$= c, \forall t$$

$$x(t) - x(t-\tau) = 0, \forall t, \tau$$

$$|x(t) - x(t-\tau)|^2 = 0, \forall t, \tau$$

$$E\{|X(t) - X(t-\tau)|^2\} = 0, \forall t, \tau$$

2  $R_{xx}(t)$

WS stationarity

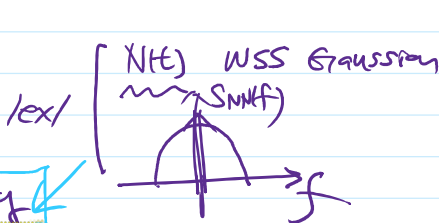
$$(i) \mu_X(t) = c, \forall t$$

$$(ii) r_{xx}(t, s) = r_{xx}(t-s, 0), \forall(t, s)$$

WS periodicity

$$(i) \mu_X(t) = \mu_X(t+\tau)$$

$$(ii) r_{xx}(t, s) = r_{xx}(t+mT, s+nT), \forall(t, s) \quad \forall m, n$$



$$\mu_X(t) = c, \forall t$$

$$r_{NN}(t, s) = r_{NN}(t-s)$$

$$x(t) = x(t-\tau), \forall t, \tau$$

$$= c, \forall t$$

$$x(t) = x(t-T), \forall t, T \in \mathbb{R}$$

periodicity

$$R_{xx}(t, t) = E\{X(t)X(t)\}$$

$$= E\{X(t-\tau)X(t-\tau)\}$$

$$R_{xx}(t, t) - R_{xx}(t, t-\tau) - R_{xx}(t-\tau, t) + R_{xx}(t-\tau, t-\tau) = 0, \forall t, \tau$$

almost surely

$$X(t) = X(t-\tau) \text{ for all } t, \tau$$

$$E\{X(t-\tau)X^*(t-\tau)\}$$

$$= E\{X(t)X^*(t)\}$$

$$= R_{xx}(t, t)$$

□ Lemma.  $X(t) = X(t+T), \forall t$  in the MS sense  $\Rightarrow$

$$(i) \mu_X(t) = \mu_X(t+T), \forall t \text{ \&}$$

$$(ii) R_{XX}(t, s) = R_{XX}(t+mT, s+mT), \forall (t, s), \forall (m, n)$$

(proof)

$$(i) 0 \leq |E\{X(t) - X(t+T)\}|^2 \leq E\{|X(t) - X(t+T)|^2\} = 0, \forall t.$$

$$\Rightarrow E\{X(t) - X(t+T)\} = 0, \forall t$$

$$\Leftrightarrow \mu_X(t) = \mu_X(t+T), \forall t$$

$$(ii) |E\{X(t)(X(s) - X(s+T))^*\}|^2 \leq E\{|X(t)|^2\} \underbrace{E\{|X(s) - X(s+T)|^2\}}_{=0, \forall s} = 0, \forall (t, s)$$

$$\Rightarrow E\{X(t)(X(s) - X(s+T))^*\} = 0, \forall (t, s)$$

$$\Leftrightarrow R_{XX}(t, s) = R_{XX}(t, s+T).$$

Similarly

$$R_{xx}(t, s) = R_{xx}(t+T, s)$$

$$\begin{aligned} \therefore R_{xx}(t, s) &= R_{xx}(t, s+T) = R_{xx}(t, s+nT) \\ &= R_{xx}(t+T, s) = R_{xx}(t+mT, s) \\ &= R_{xx}(t+mT, s+nT) \quad \forall(t, s) \quad \forall(m, n) \end{aligned}$$

□ MS stationarity ?

□ Def.  $E\{|X(t) - X(t-\tau)|^2\} = 0, \quad \forall(t, \tau)$

□ Lemma. MSS  $\Rightarrow$  WSS

(i)  $0 \leq E\{|X(t) - X(t-\tau)|^2\} \leq E\{|X(t) - X(t-\tau)|^2\} = 0, \quad \forall(t, \tau)$

$$\Rightarrow E\{X(t) - X(t-\tau)\} = 0, \quad \forall(t, \tau)$$

$$\Leftrightarrow \mu_X(t) = \mu_X(t-\tau), \quad \forall(t, \tau)$$

$$\Rightarrow \mu_X(t) = c, \quad \forall t$$

(ii)  $E\{X(t)(X(s) - X(s-\tau))^*\} = 0 \leq E\{|X(t) - X(t-\tau)|^2\} = 0, \quad \forall(t, s, \tau)$

$\Rightarrow E\{X(t)(X(s) - X(s-\tau))^*\} = 0, \quad \forall(t, s, \tau)$

$$\Rightarrow R_{xx}(t, s) = R_{xx}(t, s-\tau), \quad \forall(t, s, \tau)$$

$$\Leftrightarrow R_{xx}(t, s) = R_{xx}(t, s') \quad \forall(t, s, s')$$

Similarly

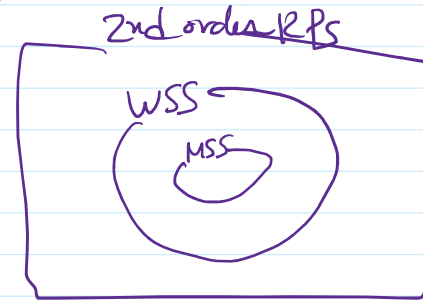
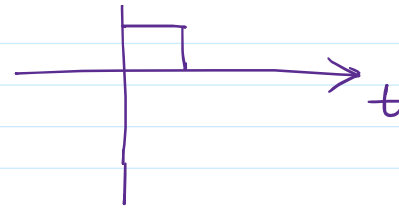


Similarly

$$R_{xx}(t, s) = R_{xx}(t', s') \quad \forall (t, s, t', s')$$

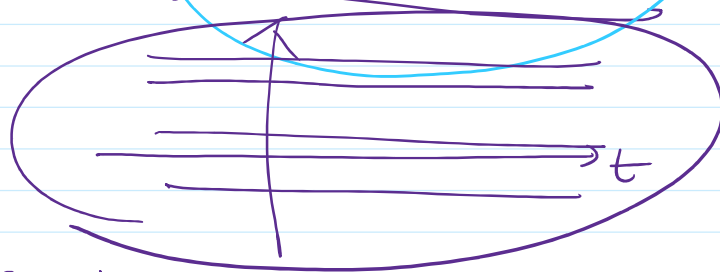
$$\Rightarrow R_{xx}(t, s) = \sigma^2 \quad \forall (t, s) \quad (1)$$

$$R_{xx}(t, s) = R_{xx}(t-s, 0) \quad \forall (t, s)$$



"No meaning!"

$X(t) = X, \forall t. E[X] = c$   
MS sense.  $E[X^2] = \sigma^2$

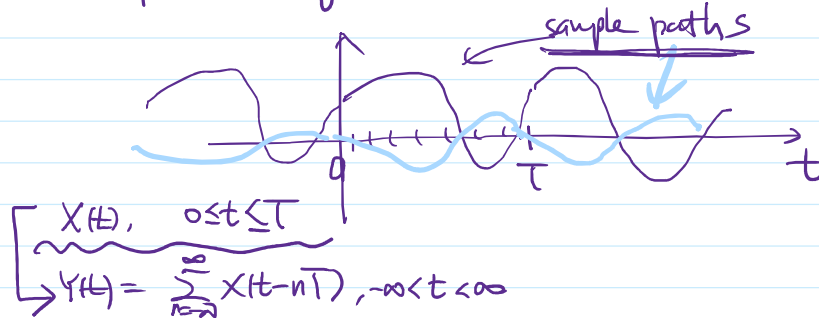


A RV is generated & repeats for all  $t$ .

→ No much meaning.



## MS Periodicity



→ Sometimes, we only have a finite-interval observation.

For  $[0, T)$ , we have uncountably infinite # of correlated RVs.

Then, this period repeats to make the sample path a periodic function

For convenience, we may periodically extend it to make it

an MS extended RD. Also, the same analysis can be done for

For convenience, we may periodically extend it to make it an MS periodic RP, where we may apply Fourier basis for analysis. For example, K-L expansion may be possible wr

Fourier basis.

$$C_{xx} \underline{v}_i = \lambda_i \underline{v}_i \quad \left( \int_0^T C_{xx}(t,s) \underline{v}_i(s) ds = \lambda_i \underline{v}_i(t), \quad 0 \leq t,s < T \right)$$

$$\Rightarrow C_{xx} = \sum \lambda_i \underline{v}_i \underline{v}_i^H \quad \Leftrightarrow \quad C_{xx}(t,s) = \sum_i \lambda_i \underline{v}_i(t) \underline{v}_i^*(s), \quad 0 \leq t,s < T$$

$V = [\underline{v}_1 \dots \underline{v}_N]^T$

$$\Rightarrow C_{xx} = V \Lambda V^H \quad \text{where } \int_0^T \underline{v}_i(t) \underline{v}_j^*(t) dt = \delta_{ij}$$

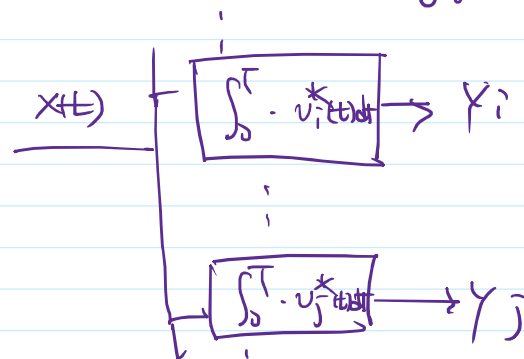
$$\Rightarrow \underline{x} \xrightarrow{V^H} \underline{Y}$$

$$\Rightarrow \underline{Y} = V^H \underline{x}$$

$$\Rightarrow \underline{x} = V \underline{Y}$$

$$\text{Cov}(\underline{Y}) = \Lambda$$

$$= \text{diag}(\Lambda) \Rightarrow$$



$$\underline{x}(t) = \sum_i Y_i \underline{v}_i(t), \quad 0 \leq t < T$$

where  $\text{Cov}(Y_i, Y_j) = \lambda_i \delta_{ij}$  i.e., uncorrelated RVs.

$v_k(t)$  can be  $e^{j2\pi \frac{k}{T} t}$ , then  $Y_k$  is the Fourier Coefficient

$$\underline{x}(t) = \sum Y_k e^{j2\pi \frac{k}{T} t}, \quad 0 \leq t < T$$

$$x(t) = \sum_k Y_k e^{j 2\pi k t / T}, \quad 0 \leq t < T.$$

$$\text{cov}(Y_i, Y_j)$$

□ A time-limited 2nd-order RP  $X(t)$ ,  $0 \leq t \leq T$ .

$E\{X(t)\} = \mu_{X(t)}$ ,  $0 \leq t \leq T$

$R^* R$ .  $E\{X(t)X^*(s)\} = r_{XX}(t,s)$ ,  $0 \leq t,s \leq T$

$= E\{X^*(s)X(t)\}$

$\int_0^T r_{XX}(t,s) v_i(s) ds = \lambda_i v_i(t)$ ,  $0 \leq t \leq T$ . (Kernel)

$\int_0^T \int_0^T a^*(t) r_{XX}(t,s) a(s) dt ds = E\left\{\left|\int_0^T a^*(t) X(t) dt\right|^2\right\} \geq 0$

$\int_0^T v_i(t) v_j^*(t) dt = \delta_{ij}$

$(\lambda_i, v_i(t))$ ; Fredholm Integral Equations

$v_i(t) \rightarrow \text{LTV Filter w/ } r_{XX}(t,s) \rightarrow \lambda_i v_i(t)$

$v_i(t) = 0, t < 0, t > T$

$E\left\{\left(\int_0^T a^*(t) X(t) dt\right) \left(\int_0^T a^*(s) X(s) ds\right)^*\right\}$

$= E\left[\int_0^T \int_0^T a^*(t) X(t) X^*(s) a(s) dt ds\right]$

$= \int_0^T \int_0^T a^*(t) r_{XX}(t,s) a(s) dt ds$

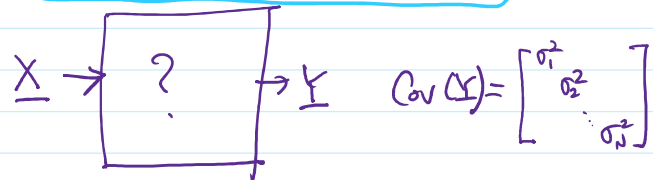
$$r_{XX}(t,s) = \sum_i \lambda_i v_i(t) v_i^*(s)$$

$$R = \sum_i \lambda_i \underline{v}_i \underline{v}_i^H$$

$$= V \Lambda V^H$$

$\Lambda = \text{diag}(\lambda_1, \lambda_2, \dots)$

$V = [\underline{v}_1, \underline{v}_2, \dots]$



$E[\underline{X}] = \underline{\mu}$

$\text{cov}(\underline{X}) = C_{XX} = C_{XX}^H \succeq 0$

$= V \Lambda V^H = \underline{V} \underline{\Lambda} \underline{V}^H$

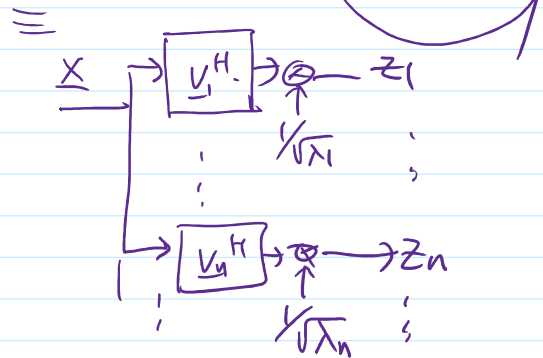
$\underline{Y} = \underline{V}^H \underline{X}$

$\underline{X} = \underline{V} \underline{Y}$

$$\underline{X} \underline{X}^H = \underline{V} \underline{\Lambda} \underline{V}^H = \underline{V} \underline{\Lambda}^{\frac{1}{2}} \underline{\Lambda}^{\frac{1}{2}} \underline{V}^H \xrightarrow{\substack{\underline{Y} = \underline{V}^H \underline{X} \\ \underline{\Lambda}^{\frac{1}{2}} \underline{\Lambda}^{\frac{1}{2}} = \underline{\Lambda}}} \underline{X} = \underline{V} \underline{Y}$$

$$\begin{aligned} \underline{X} &\rightarrow \boxed{\underline{V}^H} \rightarrow \underline{Y} \quad \underline{E}[\underline{Y}] = \underline{V}^H \underline{\mu}_X \\ &\quad \downarrow \underline{V}^{-1} \quad \text{Cov}[\underline{Y}] = \text{Cov}[\underline{V}^H \underline{X}] \\ &\quad \quad \quad = E[\underline{V}^H (\underline{X} - \underline{\mu}_X) (\underline{V}^H (\underline{X} - \underline{\mu}_X))^H] \\ &\quad \quad \quad = \underline{V}^H \text{Cov}[\underline{X}] \underline{V} = \underline{V}^H (\underline{V} \underline{\Lambda} \underline{V}^H) \underline{V} = \underline{\Lambda} \end{aligned}$$

$$\underline{X} \rightarrow \boxed{\underline{\Lambda}^{-\frac{1}{2}} \underline{V}^H} \rightarrow \underline{Z} \quad \begin{aligned} \underline{E}[\underline{Z}] &= \underline{\Lambda}^{-\frac{1}{2}} \underline{V}^H \underline{\mu}_X \\ \text{Cov}[\underline{Z}] &= \underline{I} \end{aligned} \quad \boxed{\underline{X} = \underline{V} \underline{\Lambda}^{\frac{1}{2}} \underline{Z}}$$



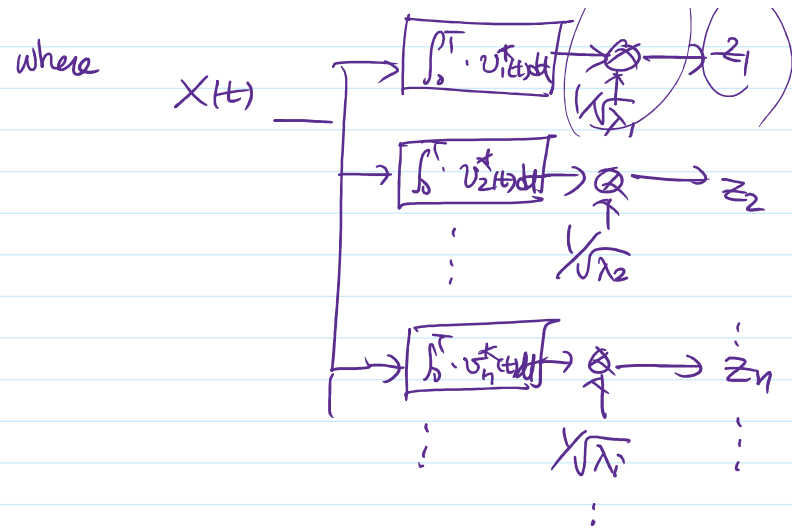
$$\begin{aligned} &= \underline{V} \begin{bmatrix} \sqrt{\lambda_1} z_1 \\ \vdots \\ \sqrt{\lambda_n} z_n \end{bmatrix} = [\underline{v}_1 \underline{v}_2 \dots \underline{v}_n] \begin{bmatrix} \sqrt{\lambda_1} z_1 \\ \vdots \\ \sqrt{\lambda_n} z_n \end{bmatrix} \\ &= \sum_i \sqrt{\lambda_i} z_i \underline{v}_i \end{aligned}$$

$z_i$ 's are uncorrelated RVs w/ same variance limit.

$$\begin{aligned} \left\{ \begin{array}{l} \underline{Z} \sim \underline{\mu}_Z \\ \text{Cov}(\underline{Z}) = \underline{I} \end{array} \right. & \quad \begin{array}{c} z_1 \\ z_2 \\ \vdots \\ z_n \end{array} \quad \begin{array}{c} \sqrt{\lambda_1} \\ \sqrt{\lambda_2} \\ \vdots \\ \sqrt{\lambda_n} \end{array} \\ & \quad \begin{array}{c} v_1(t) \\ v_2(t) \\ \vdots \\ v_n(t) \end{array} \rightarrow \bigcirc \rightarrow \underline{X}(t) = \sum_i \sqrt{\lambda_i} z_i v_i(t) \end{aligned}$$

$\int v_i(t) v_j^*(t) dt = \delta_{ij}$

where  $\underline{X}(t) \rightarrow \boxed{\int_0^T v_i(t) dt} \rightarrow z_i$



SIS and Nyquist Criterion, Orthogonality vs. Bi-orthogonality,  
SNR, SIR, SINR, Zero-Forcing vs. MMSE

2025년 11월 5일 수요일  
10:26

## II Nyquist Criterion for zero ISI

↑ intersymbol interference

Def. if  $\int_{-\infty}^{\infty} p(t-mT) p^*(t-nT) dt = C \delta_{mn}$ ,  $C > 0$ ,  $\forall m, n \in \mathbb{Z}$ .

$p(t)$  satisfies a Nyquist criterion for T.

when the symbol rate is  $1/T$ .

$\left( \begin{array}{l} m=n \Rightarrow \int_{-\infty}^{\infty} |p(t)|^2 dt = C \\ m \neq n \Rightarrow \int_{-\infty}^{\infty} p(t) p^*(t-nT) dt = 0 \end{array} \right)$

Caution!

RP

Caution!

$$X(t) = \sum_{n=-\infty}^{\infty} d_n p(t-nT) \quad \leftarrow \text{RP}$$

$$x(t) = \sum_{n=-\infty}^{\infty} a_n p(t-nT)$$

(?)  $(d_n)_n$  is a data symbol seq. w/  $d_n$ 's are RVs.

(?)  $(a_n)_n$  " a deterministic sequence w/ finite energy.

